

# The Heston Stochastic-Local Volatility Model: Efficient Monte Carlo Simulation

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## 1 Introduction

The following method proposed by [1] deals with an efficient Monte Carlo scheme for simulating the stochastic volatility model of Heston enhanced by a non-parametric local volatility component.

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## 2 Theoretical framework

We consider the following local stochastic volatility model

$$\begin{aligned}\frac{dS_t}{S_t} &= r dt + \sigma(t, S_t)\psi(V_t)dW_t^x, \\ dV_t &= a_v(t, V_t)dt + b_v(t, V_t)dW_t^v, \\ dW_t^x dW_t^v &= \rho_{x,v}dt,\end{aligned}$$

where  $r$  denotes the risk-free interest rate,  $\rho_{x,v}$  is the correlation between the corresponding Brownian motions,  $\sigma(t, S_t)$  is the local volatility component,  $\psi(V_t)$  controls the stochastic volatility, parameters  $a_v(t, V_t)$  and  $b_v(t, V_t)$  determine the drift and diffusion of the variance process respectively. In the following, we consider

$$\begin{aligned}\psi(V_t) &= \sqrt{V_t}, \\ a_v(t, V_t) &= \kappa(\bar{v} - V_t), \\ b_v(t, V_t) &= \gamma\sqrt{V_t},\end{aligned}$$

where  $\kappa$  controls the speed of mean reversion,  $\bar{v}$  controls a long-term mean and  $\gamma$  determines the volatility of the process  $V_t$ .

To be able to calibrate market smiles exactly, the volatility  $\sigma$  is given by the following equality

$$\sigma^2(t, K) = \frac{\sigma_{LV}^2(t, K)}{\mathbb{E}[\psi^2(V_t)|S_t = K]}$$

where  $\sigma_{LV}$  denotes the Dupire's local volatility.

### 3 Numerical algorithm

#### 3.1 Computation of $\mathbb{E}[\psi^2(V_t)|S_t = K]$

Suppose that at a given time  $t_i$ ,  $i = 1, \dots, N$  we have  $M$  pairs of Monte Carlo realizations  $(s_{i,1}, v_{i,1}), \dots, (s_{i,M}, v_{i,M})$ . By grouping the pairs of realizations into bundles  $[b_{i,1}, b_{i,2}], [b_{i,2}, b_{i,3}], \dots, [b_{i,l}, b_{i,l+1}]$ , we get

$$\begin{aligned} \mathbb{E}[\psi^2(V_{t_i})|S_{t_i} = s_{i,j}] &\sim \mathbb{E}[\psi^2(V_{t_i})|S_{t_i} \in ]b_{i,k}, b_{i,k+1}[], \\ &\sim \frac{\mathbb{E}[\psi^2(V_{t_i})\mathbb{1}_{\{S_{t_i} \in ]b_{i,k}, b_{i,k+1}[}\}}{\mathbb{P}(S_{t_i} \in ]b_{i,k}, b_{i,k+1}[}) \end{aligned}$$

for  $s_{i,j} \in ]b_{i,k}, b_{i,k+1}[$ .

If we build the  $l$  bins such that each bin contains the same number of Monte Carlo paths, we get

$$\begin{aligned} \mathbb{E}[\psi^2(V_{t_i})|S_{t_i} = s_{i,j}] &\sim \frac{\frac{1}{M} \sum_{j=1}^M \psi^2(v_{i,j}) \mathbb{1}_{\{s_{i,j} \in ]b_{i,k}, b_{i,k+1}[}\}}{\mathbb{P}(S_{t_i} \in ]b_{i,k}, b_{i,k+1}[}) \\ &\sim \frac{l}{M} \sum_{j \in \mathcal{J}_{i,k}} \psi^2(v_{i,j}) \end{aligned}$$

where  $\frac{1}{l}$  represents the probability to be in the  $k$ th bin and  $\mathcal{J}_{i,k} := \{j | s_{i,j} \in ]b_{i,k}, b_{i,k+1}[}\}$ . We summarize the method:

For each step  $t_i$ ,  $i = 1, \dots, N$

1. Generate  $M$  pairs of observations  $(s_{i,j}, \psi^2(v_{i,j}))$ ,  $j = 1, \dots, M$ .
2. Order the elements  $\bar{s}_{i,j} : \bar{s}_{i,1} \leq \dots \leq \bar{s}_{i,M}$  and apply the same permutation on  $(v_{i,1}, \dots, v_{i,M})$ .
3. Determine the boundary of the  $l$  bins  $]b_{i,k}, b_{i,k+1}[$ ,  $k = 1, \dots, l$  in the following way

$$b_{i,1} = \bar{s}_{i,1}, \quad b_{i,l+1} = \bar{s}_{i,N}, \quad b_{i,k} = \bar{s}_{i,(k-1)M/l}, \quad k = 2, \dots, l.$$

4. For the  $k$ th bin approximate the conditional expectation by

$$\mathbb{E}[\psi^2(V_{t_i})|S_{t_i} \in ]b_{i,k}, b_{i,k+1}[] \sim \frac{l}{M} \sum_{j \in \mathcal{J}_{i,k}} \psi^2(v_{i,j})$$

### 3.2 Simulation scheme

First, we recall the dynamics of the Heston SLV model expressed in terms of independent Brownian motions:

$$\begin{aligned}\frac{dS_t}{S_t} &= r dt + \sigma(t, S_t) \sqrt{V_t} \left( \rho_{x,v} d\tilde{W}_t^v + \sqrt{1 - \rho_{x,v}^2} d\tilde{W}_t^x \right), \\ dV_t &= \kappa(\bar{v} - V_t)dt + \gamma \sqrt{V_t} d\tilde{W}_t^v\end{aligned}$$

where  $\tilde{W}^x$  and  $\tilde{W}^v$  are independent Brownian motions. Following the scheme proposed in [1, Section 3.3], we discretize  $[0, T]$  on a regular grid of size  $N$ , with step size  $\Delta = \frac{T}{N}$ . We get

$$\begin{aligned}v_{i+1,j} &\sim c(\Delta) \chi^2(d, \lambda(t_i, v_{i,j})), \\ x_{i+1,j} &= x_{i,j} + r\Delta - \frac{1}{2} \hat{\sigma}^2(t_i, x_{i,j}) v_{i,j} \Delta + \frac{\rho_{x,v}}{\gamma} \hat{\sigma}(t_i, x_{i,j}) (v_{i+1,j} - \kappa \bar{v} \Delta + v_{i,j} c_1) \\ &\quad + \rho_1 \sqrt{\hat{\sigma}^2(t_i, x_{i,j}) v_{i,j} \Delta} \tilde{Z}_x,\end{aligned}$$

where  $x_{i,j} = \ln(s_{i,j})$ ,  $\rho_1 = (1 - \rho_{x,v}^2)^{1/2}$ ,  $c_1 = \kappa \Delta - 1$  and

$$c(\Delta) = \frac{\gamma^2}{4\kappa} (1 - e^{-\kappa \Delta}), \quad d = \frac{4\kappa \bar{v}}{\gamma^2}, \quad \lambda(t, x) = \frac{4\kappa e^{-\kappa \Delta}}{\gamma^2 (1 - e^{-\kappa \Delta})} x,$$

$\chi^2(d, x)$  represents a noncentral chi-squared distribution with  $d$  degrees of freedom and non-centrality parameter  $x$ , and

$$\hat{\sigma}^2(t_i, x_{i,j}) = \sigma^2(t_i, e^{x_{i,j}}) = \frac{\sigma_{LV}^2(t_i, s_{i,j})}{\mathbb{E}[V_{t_i} | S_{t_i} = s_{i,j}]}.$$

## 4 Numerical experiments

We test the algorithm on a Call option with payoff  $e^{-rT}(S_T - K)_+$  with the following parameters, for different maturities and strikes :

$r$	$s_0$	$v_0$	$\kappa$	$\gamma$	$\rho_{x,v}$
0	1	0.0945	1.05	0.95	-0.315

We use  $l = 20$  bins,  $N = 100$  times steps and  $M = 10^5$  Monte Carlo simulations and the local volatility  $\sigma_{LV}$  is given by

$$\sigma_{LV}(t, x) = 0.01 + 0.1e^{-x/s_0} + 0.01t.$$

K T	2	8
0.7	0.300028	0.309048
0.9	0.103869	0.154880
1.1	0.004426	0.059615
1.3	0.000046	0.019109
1.5	0.000002	0.005046

## References

- [1] A.W. Van Der Stoep, L.A. Grzelak, C.W. Oosterlee. The Heston Stochastic-Local Volatility Model: Efficient Monte Carlo Simulation. 2013.

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